

Today
 - bit more on random walks
 on graphs
 - Monte Carlo methods
 & approx counting

Random walks on undirected graphs

$G = (V, E)$ undirected graph

consider simple random walk on this graph:

$$p_i = \frac{1}{d_i} \quad \forall (i,j) \in E$$

[MC is periodic iff graph is bipartite]
 [but if so, consider lazy r.w.]

$$\pi_i = \frac{d_i}{2m} \quad m: \# \text{ of edges}$$

$$\sum_i \frac{d_i}{2m} = 1$$

$$\pi_j = \sum_i \pi_i p_{ij}$$

$$\frac{d_j}{2m} = \sum_{\substack{\text{inst.} \\ (i,j) \in E}} \frac{d_i}{2m} \cdot \frac{1}{d_i}$$

Some key quantities:

hitting time

$$h_{ij} = E(T_{ij})$$

commute time

$$c_{ij} = h_{ij} + h_{ji}$$

covetime

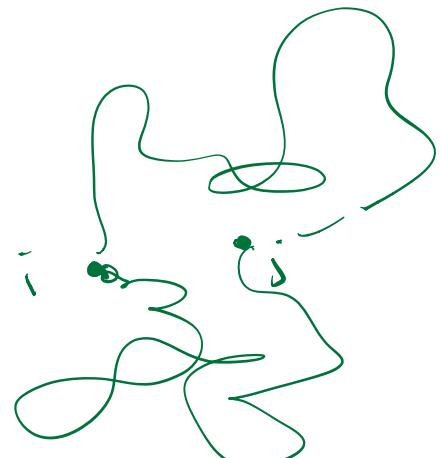
$$C(G) = \underset{\text{9x9}}{\text{time to visit all vertices}}$$

$$C(G)$$

$$h_{ii} = \frac{1}{\pi_i} = \frac{2m}{d_i}$$

Lemma

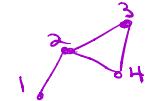
$$\forall \text{ edge } (i,j) \quad h_{ij} + h_{ji} \leq 2m$$



Pf.: Consider corresponding random walk
on directed edges (states are $\overset{i \rightarrow j}{j}$ if $(i,j) \in E$) $2m$ states

transition probabilities:

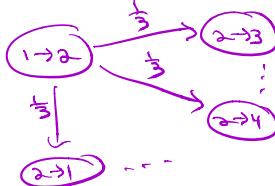
$$q_{(i \rightarrow j, l \rightarrow r)} = \begin{cases} \frac{1}{d_j} & l=j \\ 0 & \text{o.w.} \end{cases}$$



$$Q = (q_{\vec{i}, \vec{j}}) \quad \text{is doubly stochastic}$$

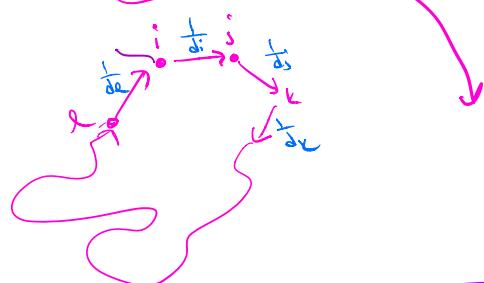
$$\text{col sum for } i \rightarrow j = \sum_{\substack{k \text{ s.t.} \\ (k, i) \in E}} q_{k \rightarrow i, i \rightarrow j} = \sum_{\substack{k \text{ s.t.} \\ (k, i) \in E}} \frac{1}{d_i} = 1$$

state attract



$\Rightarrow \pi$ is uniform, i.e. $\pi_{\vec{i}} = \frac{1}{2m} \forall \vec{i}$

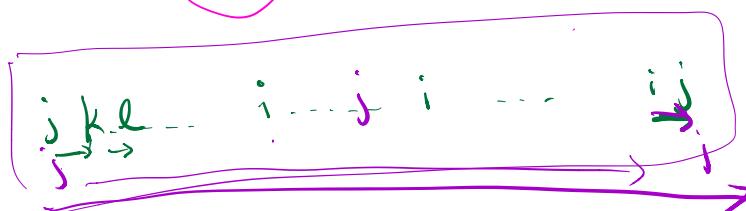
$$\Rightarrow h_{i \rightarrow j, i \rightarrow j} = 2m$$



$$h_{i \rightarrow j, i \rightarrow j} = 2m$$

$\underbrace{h_{ij} + h_{ji}}$
original r.w.
 $\underbrace{\text{in other walks}}$

① $\xleftarrow{\dots} i \rightarrow j$



Corollary

If $G = (V, E)$
connected
nonbipartite
(or lazy)
 $C(G) = \text{expected cover time}$
of random walk
on $G \leq \underline{2m(n-1)}$

$O(mn)$

Pf: Let T be a spanning tree on G
and let $e_{ij}, e_{ji} \in \{e_{ij}, e_{ji}\}$ be the edges in the tree.

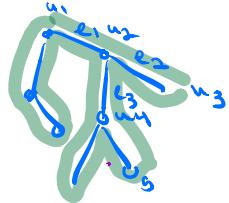
Consider doubled tree T^* , where each edge is duplicated once in each direction.

Every vertex has indegree = outdegree
 \Rightarrow has Euler tour say

$$\overrightarrow{e_1} \overrightarrow{e_2} \dots \overrightarrow{e_i} \dots$$

$$v_1 v_2 \dots v_i \dots$$

$$E(\text{cover time}) \leq \sum_{i=1}^m E(T_{u_iv_i} + T_{v_iu_i}) = (n-1)2m$$

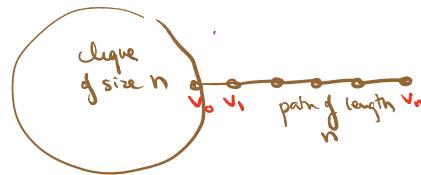


$$u_1 u_2 u_3 u_2 u_4 u_5 u_6 \dots u_7 u_2$$

Examples



$$C(G) = \Theta(n^2)$$



$$m = O(n^2)$$

$$C(G) = O(n^3)$$

Fact: Starting from v_1 , $\Pr(\text{reach } v_n \text{ before returning to } v_0) = \frac{n-1}{n}$

$$\frac{n-1}{n}$$

$$\text{bound above } C(G) = O(n^3)$$



$$O(n \log n) - \text{coupon collectors}$$

$$\frac{n-i}{n} \quad E(T \text{ to see new}) = \frac{n}{n-i}$$

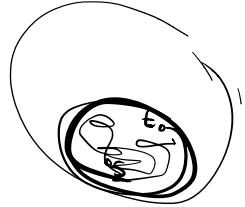
#balls \rightarrow n bins fill every bin has a ball.

Applications: s-t connectivity

Given G undirected graph $s, t \in V$
determine if s & t are in same CC.

DFS $\begin{cases} O(m) \text{ time} \\ O(n) \text{ space} \end{cases}$ keep track of all vertices
BFS $\begin{cases} O(n) \text{ time} \\ O(n) \text{ space} \end{cases}$ search has visited so far

Observation: very simple randomized alg using log space
(input on separate read-only tape)



Simulate r.w. of length Hmn on G starting from s

$$\Pr(\text{r.w. doesn't reach } t \text{ when } \exists \text{ path}) \leq \frac{1}{2}$$

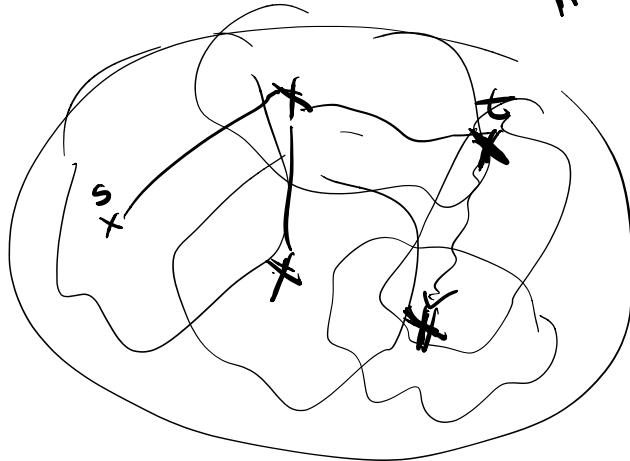
Both algorithms $\tilde{O}(mn)$ space-time
logarithmic factors being ignored

Can we interpolate?

$\tilde{O}(p)$ space

$\tilde{O}\left(\frac{mn}{p}\right)$ time?

Feige



random walks \longleftrightarrow electrical networks

Monte Carlo Methods

collection of tools for estimating values thru sampling & estimation

(ϵ, δ) Approximation

A randomized alg gives an (ϵ, δ) approx for value V if the output X of the alg satisfies

$$\Pr(|X - V| > \epsilon |V|) \leq \delta$$

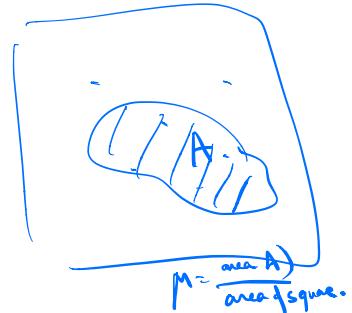
Example

Sample indep random vars whose mean is quantity we want to estimate

Let X_1, X_2, \dots, X_m iid. Bernoulli with $E(X_i) = \mu$

$$\text{If } m \geq \frac{3 \ln(\frac{2}{\delta})}{\epsilon^2 \mu} \text{ then}$$

$$\Pr\left(\left|\frac{1}{m} \sum_{i=1}^m X_i - \mu\right| > \epsilon \mu\right) \leq \delta \quad \text{Pf: Chernoff bounds}$$



Monte Carlo Thm

DNF Counting

Suppose want to know # satisfying assignments

$$(\bar{x}_1 \wedge x_2) \vee (x_2 \wedge x_3) \vee (x_1 \wedge \bar{x}_2 \wedge \bar{x}_3 \wedge x_4) \vee (x_3 \wedge \bar{x}_4)$$

Obviously satisfying such a formula is easy

Counting # satisfying assignments is hard

If could do this, could solve ~~X~~SAT

CNF formula $\varphi \xrightarrow{\text{DNNF}} \overline{\varphi}$

If count = 2^n
from φ unsat
o.w. φ sat.

DNF $\overline{\varphi}$
↓
count # solns

$$(x_1 \vee x_2 \vee \bar{x}_3) \wedge \dots$$

$$(\bar{x}_1 \wedge \bar{x}_2 \wedge x_3) \wedge \dots$$

Problem actually #P-complete \rightarrow strong intractability

#P is counting analogue of NP

problems of form:
compute $f(x)$ where $f(x)$
is # solutions to problem x in NP

counting # sat assignments DNF formula #P complete
Ham. cycles

matchings in bipartite graph.

P=NP

Approximate DNF Counting?

Obvious approach: sample random assignments, indep m times
 $X_i = \begin{cases} 1 & \text{if random assignment satisfies } \psi \\ 0 & \text{else.} \end{cases}$

return $\frac{\sum X_i}{m} \cdot 2^n$ as estimate for # satisfying assignments

for (ϵ, δ) estimate need $m = \Omega\left(\frac{\ln(\frac{2}{\delta})}{\epsilon^2}\right)$

m could be exponentially small $\frac{n^2}{2^m}$ fraction of assign.

finding needle in a haystack $n \approx n^2 \approx n^3$

3 FPRAS for DNF counting

Fully polynomial randomized approx scheme

aka: a randomized alg. for which, given an input x and any parameters ϵ, δ with $0 < \epsilon, \delta < 1$, the alg. outputs an (ϵ, δ) approx to $V(x)$ in time poly in $\frac{1}{\epsilon}, \ln \frac{1}{\delta}$ & size of input

in this example
 x is DNF formula
 $V(x)$ is # satisfying assignments

$$\bar{V} \text{ is } (\epsilon, \delta) \text{ approx to } V \equiv \Pr((\bar{V} - V) > \epsilon V) \leq \delta$$

$$\psi = C_1 \vee C_2 \vee \dots \vee C_t$$

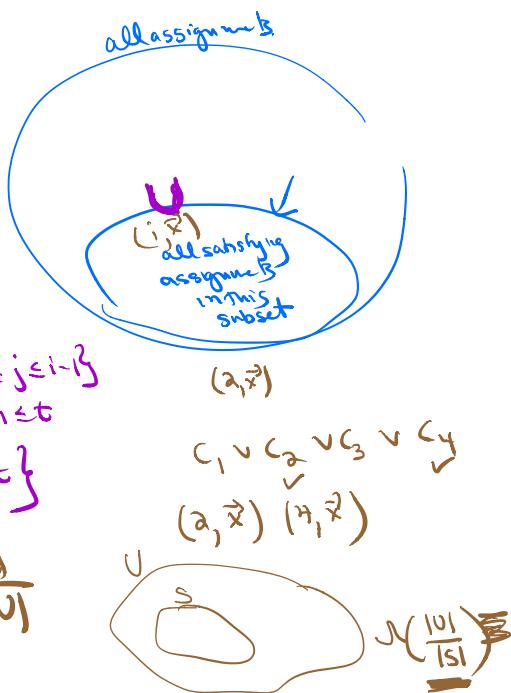
SC_i : set of assignments satisfying clause i

S = set of all satisfying assignments

$$= \{(i, \vec{x}) \mid \vec{x} \in SC_i, \vec{x} \notin SC_j, \substack{i \leq j \leq t \\ 1 \leq i \leq t}\}$$

$$U = \{(i, \vec{x}) \mid \vec{x} \in SC_i, 1 \leq i \leq t\}$$

estimate $|S|$ by approximately $\frac{|S|}{|U|}$



M times, sample (i, x) from U & then check
if its in S

$$|U| = \sum_{i=1}^t |SC_i|$$

#rows not in C_i

$$\frac{|S|}{|U|} \geq \frac{1}{t}$$

$$M = \frac{3 + \ln(\frac{\alpha}{\delta})}{\epsilon^2} \text{ samples}$$

pick i with prob $\frac{|SC_i|}{\sum_j |SC_j|}$ (ϵ, δ) approx

$$\Pr((i, x) \text{ selected}) = \frac{\text{sets rows outside } SC_i}{\sum_j |SC_j|} \stackrel{\text{u.a.r}}{=} \frac{1}{|U|}$$

DNF counting illustrates fundamental connection
 Sampling \leftrightarrow counting

From approx sampling \rightarrow approx counting

$$\begin{array}{l} S \subseteq V \\ \text{Is. } S \text{ I.S.} \\ \forall (i,j) \in S \quad (i,j) \notin E \end{array}$$

Counting Independent Sets

$$\begin{array}{l} \text{Input:} \\ X, \epsilon \\ \uparrow \\ G = (V, E) \end{array}$$

FPAUS
 Fully polynomial
 almost uniform
 sampler

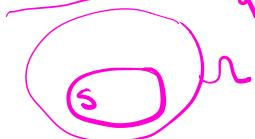
runs in time
 polynomial in size of X
 $\& \ln(\frac{1}{\epsilon})$

input $G = (V, E)$
 output: estimate of

$$|\mathcal{I}(G)|$$

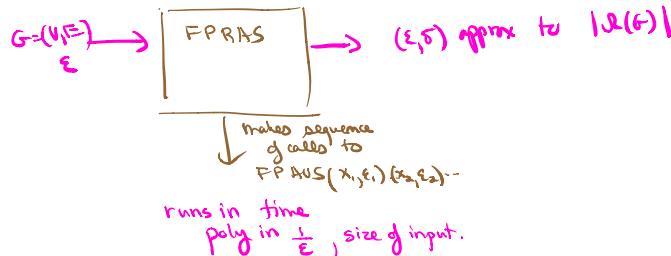
independent sets
 in graph.

$$\epsilon \text{ uniform sample } w \text{ from } \mathcal{I}(x) \quad |\Pr(w \in S) - \frac{|S|}{|\mathcal{I}(x)|}| \leq \epsilon$$



Want to show:

Given FPAUS for indeps \Rightarrow can construct FPRAS
 for ISSs



$$G = (V, E) \quad e_1, e_2, \dots, e_m \text{ arbitrary ordering of edges}$$

$$E_i = \{e_1, e_2, \dots, e_i\} \quad G_i = (V, E_i) \quad G_m = G \quad G_0 = (V, \emptyset)$$

$\mathcal{I}(G_i)$: set of ISSs in G_i

$$|\mathcal{I}(G)| = \frac{|\mathcal{I}(G_m)|}{|\mathcal{I}(G_{m-1})|} \times \frac{|\mathcal{I}(G_{m-1})|}{|\mathcal{I}(G_{m-2})|} \times \dots \times \frac{|\mathcal{I}(G_1)|}{|\mathcal{I}(G_0)|} \times 2^n$$

$$\text{Let } r_i = \frac{|\mathcal{I}(G_i)|}{|\mathcal{I}(G_{i-1})|}$$

$$|\mathcal{I}(G)| = r_m \cdot r_{m-1} \cdots r_1 \cdot 2^n$$

approx r_i by \tilde{r}_i and output as estimate for $|\mathcal{I}(G)|$

$$X = 2^n \prod_{i=1}^m \tilde{r}_i$$

So need to bound $R = \prod_{i=1}^m \tilde{r}_i^{n_i}$

Claim: If \tilde{r}_i is $(\frac{\epsilon}{2m}, \frac{\delta}{m})$ -approx to r_i $\forall i$
 then X is (ϵ, δ) -approx to $|JL(G)|$

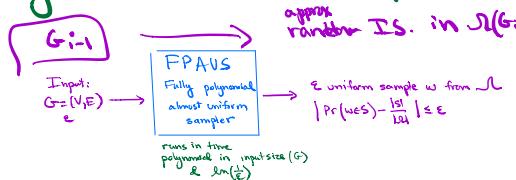
$$\begin{aligned} \text{Pf: } \Pr(|\tilde{r}_i - r_i| \leq \frac{\epsilon}{2m} r_i) &\geq 1 - \frac{\delta}{m} \quad \forall i \\ \Rightarrow r_i(1 - \frac{\epsilon}{2m}) \leq \tilde{r}_i \leq r_i(1 + \frac{\epsilon}{2m}) &\quad \text{w.p.} \geq 1 - \frac{\delta}{m} \\ \Rightarrow (1 - \frac{\epsilon}{2m}) \leq \frac{\tilde{r}_i}{r_i} \leq (1 + \frac{\epsilon}{2m}) & \\ 1 - \epsilon \leq (1 - \frac{\epsilon}{2m})^m \leq \prod_{i=1}^m \frac{\tilde{r}_i}{r_i} \leq (1 + \frac{\epsilon}{2m})^m \leq 1 + \epsilon &\quad \text{w.p.} \geq 1 - \delta \\ (1 - \epsilon) \frac{m}{\prod_{i=1}^m r_i} \leq \frac{n}{\prod_{i=1}^m r_i} \leq (1 + \epsilon) \frac{m}{\prod_{i=1}^m r_i} & \quad \text{w.p.} 1 - \delta \\ \equiv \Pr\left(\left|\frac{\prod_{i=1}^m \tilde{r}_i}{m} - \frac{m}{\prod_{i=1}^m r_i}\right| \geq \epsilon \frac{m}{\prod_{i=1}^m r_i}\right) \leq \delta & \end{aligned}$$

To get \tilde{r}_i $(\frac{\epsilon}{2m}, \frac{\delta}{m})$ -approx for r_i ,

use FPAUS for ISS

Idea: (approx) sample indep sets in G_{i-1} &
 compute fraction of these that are indep in G_i

$$r_i = \frac{|I_e(G_i)|}{|JL(G_{i-1})|}$$

For this to work, need $r_i = \frac{|I_e(G_i)|}{|JL(G_{i-1})|}$ not too small (no needle in haystack)

Claim: $r_i \geq \frac{1}{2}$ G_i contains one extra edge (u, v)

so any Σ in G_{i-1} but not indep in G_i
 contains both u & v

$$\begin{aligned} \exists u \quad \forall \Sigma \in G_{i-1} \setminus G_i \rightarrow \Sigma \cup v \text{ in } G_i \cap G_{i-1} \\ |I_e(G_i \setminus G_i)| \leq |I_e(G_i)| \\ r_i = \frac{|I_e(G_i)|}{|JL(G_{i-1})|} = \frac{|I_e(G_i)|}{|I_e(G_i)| + |I_e(G_i \setminus G_i)|} \geq \frac{1}{2} \end{aligned}$$

\Rightarrow with polynomially many calls to FPAUS, we can get
 good approx \tilde{r}_i to r_i using Monte Carlo Thm

Two errors we need to bound:

① FPAUS \neq exact sampler so avg value $\neq r_i$

But using, say, $\frac{\epsilon}{6m}$ sampler $|E(F_i) - r_i| \leq \frac{\epsilon}{6m}$

② With samples we get approx to $E(\tilde{F}_i)$

since r_i big ($\geq \frac{1}{2}$), $E(F_i)$ big \Rightarrow

Monte Carlo Thm $\Rightarrow O\left(\frac{m^2}{\epsilon^2} \ln\left(\frac{1}{\delta}\right)\right)$ samples suffice

Thm: Given a FPAUS for sampling ISSs, one can construct a FPRAS for ISSs

Approach works for many "self-reducible" problems

Another example: counting # matchings in a graph

Again: $E = \{e_1, \dots, e_m\}$
 $G_i = (V, E_i)$ where $E_i = \{e_1, e_i\}$

$$|M(G)| = \frac{|M(G_m)|}{|M(G_{m-1})|} \cdot \frac{|M(G_{m-1})|}{|M(G_{m-2})|} \cdots \frac{|M(G_2)|}{|M(G_1)|} |M(G_1)|$$

$$\text{Like before: } \frac{|M(G_i)|}{|M(G_{i-1})|} \geq \frac{1}{2}$$

Big question remains: how to construct approx sampler?

